

Household Debt, Financialization, and Macroeconomic Performance in the U.S., 1951-2009

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Published: *Journal of Post Keynesian Economics* 35(4), pages 675-694, Summer 2013

Abstract

We empirically examine the relationship between U.S. output and household debt. To account for structural change due to financial liberalization, we divide the sample at the fourth quarter of 1982. We find structural differences between earlier and later business cycles for the U.S. household sector and its relation to the macroeconomy. In the regression analysis for pre-1982, we find no evidence that household debt variables had any negative effect on output. However, we find some evidence that household debt variables have negative effects on output for the post-1982 period. A formal structural break test provides evidence of a structural change in the relationship of U.S. output to household debt. Unit root tests for the separate samples show that none of the household variables possesses a unit root in the earlier period, yet all of them do in the later period, indicating fundamental differences between earlier and later periods in terms of the data generating process.

Key words: household debt, output, financial liberalization, structural break

JEL classifications: E29, B59

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[†]I would like to thank Robert Blecker for insightful comments and detailed suggestions. Alan Isaac and Martha Starr also offered many helpful comments. I also would like to thank referee for the helpful suggestion to improve this paper. All remaining errors are mine.

1 Introduction

The U.S. experienced a significant increase in household debt prior to the outbreak of financial crisis in 2007. Household debt outstanding as a share of GDP, for example, increased from about 45 percent in 1975 to nearly 100 percent in 2006 (see figure 1). The household debt burden increased as well. Figure 2 depicts two measures of the debt service burden: household financial obligations as a percent of disposable personal income and household debt service payments as a percent of disposable personal income. These series have been considered important debt burden measures and are used by the Federal Reserve as primary measures of the household debt burden (Greenspan, 2004).¹ Both measures also show upward trends, indicating that households' financial positions have continuously been worsened.

[Figure 1 about here.]

[Figure 2 about here.]

Cynamon and Fazzari (2008) provide an informative discussion of this accumulation of household debt from the perspective of Hyman Minsky's financial instability hypothesis.² They point out that, while household debt accumulated, household expenditure increased considerably as well. For example, the ratio of personal outlays to disposable income has increased from about 88 percent in the early 1980s to nearly 100 percent in 2007. Cynamon and Fazzari argue that, although debt-financed household expenditure provided a substantial macroeconomic stimulus between the 1980s and early 2000s, the unprecedented rise in household debt could have planted the seeds for financial instability and a serious economic downturn as indeed later occurred.

However, empirical studies of the impact of household debt on macroeconomic performance have been scarce. Palley (1994) was a pioneering empirical study that analyzed household debt and business cycles from a heterodox perspective—specifically, from the perspective of Minsky's financial instability hypothesis. Palley found that an increase in debt

(new borrowing) raised real gross national product (GNP), and an increase in the debt service burden reduced GNP, based on an autoregressive distributed lag (ADL) model.³

Palley's unstructured vector autoregression (VAR) model of changes in consumer debt, consumer debt burden, and real GNP shows that a shock to the change in consumer debt or the consumer debt burden generates an initial positive and subsequent negative GNP response respectively, both followed by a cyclical and damped response.⁴ Based on these results, Palley emphasized consumer debt and the debt burden as sources of cyclical variations.⁵

Palley's study provides a point of departure for our empirical study. We extend and improve upon the work of Palley in several respects. To account for the period of financial liberalization, we will test for a structural break in the relationship of household debt to aggregate output between pre-1980 and post-1980 periods in the U.S. macroeconomy. Based on these tests, we estimate the relationship of household debt to aggregate output separately for pre-1982 and post-1982 in the U.S. macroeconomy. We also perform unit root tests, which are an important diagnostic of data for a time-series analysis. (This diagnostic is absent in Palley's work.) More broad measures of household indebtedness are also used in this extended study. Our data span is longer, and we use GDP, which is the main measure of economic output used today, instead of GNP as used by Palley.⁶

2 Empirical Analysis

In this section, we study the empirical relationship between the level of output and measures of household debt in the U.S. economy. The household and consumer debt variables are from the Flow of Funds Accounts of the United States, and are deflated by the personal consumption price index from the Bureau of Economic Analysis (BEA). Output is measured by real GDP. We use real fixed private investment for the investment variable. Both series are from BEA.⁷ (See the data appendix for the further information.)

First, we tested household debt, consumer debt, household net worth, output, and investment for unit roots using augmented Dickey-Fuller (ADF) test statistics. We specified a constant and linear time trend, and lag lengths were determined by the Schwarz Information Criterion (SIC). According to the tests, all household variables, investment, and output have unit roots for the entire sample period, 1951Q4-2009Q1.

We use an ADL model in which the right-hand side variables are entered in both level and first differenced form in our regression specification.⁸ We utilize this empirical approach since ADL models are known to be robust to many estimation problems related to non-stationary variables (e.g., spurious regression results) (Hamilton, 1994, pp. 561-562). All the explanatory variables are lagged to prevent problems of simultaneity and reverse causality. The lagged terms are limited to one period (i.e., $t-1$) since additional lags of the explanatory variables increase multicollinearity problems and complicate the estimation.⁹

The dependent variable is the level of output (GDP). The baseline model is the following:

$$\begin{aligned}
 output &= \beta_0 + \beta_1 output_{t-1} + \beta_3 networth_{t-1} + \beta_4 \Delta networth_{t-1} & (1) \\
 &+ \beta_5 householddebt_{t-1} + \beta_6 \Delta householddebt_{t-1} + \beta_7 consumerdebt_{t-1} \\
 &+ \beta_8 \Delta consumerdebt_{t-1} + \varepsilon_t
 \end{aligned}$$

The real debt burden is proxied by the level of debt accumulation. The change in debt stock represents the flow of new net borrowing, which should provide an additional source of finance for household expenditure aside from current income and accumulated wealth.¹⁰ Therefore, our hypotheses are that the change in household debt has a positive effect but the level of debt has a negative effect on output, so $\beta_6, \beta_8 > 0$ and $\beta_5, \beta_7 < 0$. A main channel through which debt can influence GDP is the balance sheet effect via consumption. To isolate the effect of household debt on output, we therefore control for household net worth. We hypothesize that both the level and change in net worth have positive effects on output so $\beta_3, \beta_4 > 0$. Table 1 reports the regression results for the entire sample period

(1951Q4-2009Q1). We utilize the Godfrey–Breusch Lagrange Multiplier (LM) test for serial correlation in the residuals. We also utilize the ARCH LM test for autoregressive conditional heteroscedasticity (i.e., volatility clustering) in the residuals. The LM test for serial correlation is done with a two-period lag specification, and a one-period lag specification is used for the ARCH test. Model 1 is the baseline specification, which is used to narrow down the number of parameters based on the t -values of the coefficients. After narrowing down the variables using model 1, model 2 reports regressions with levels and first differences of household debt and net worth. This is intuitively plausible since consumer debt is only one component of household debt.

[Table 1 about here.]

In model 2, we can see that all the variables have the expected signs and are statistically significant at conventional levels. There is, however, the possibility of serial correlation according to the Godfrey–Breusch LM test for two lags at the 5 percent significance level. In model 3, we controlled for the level and first difference in investment. We see that this corrects for the serial correlation problem, and the qualitative results of model 2 are mostly preserved. However, the level of investment has a negative coefficient, which is counterintuitive.¹¹

2.1 Structural Breaks and Unit Roots

Our sample includes the period of “neoliberal revolution” and financial liberalization starting in the late 1970s-early 1980s (Stockhammer, 2004). In the heterodox economics literature, this era is often referred to as the beginning period of “financialization.”¹² To test for this source of structural change, we have checked the stability of regression models 2 and 3 using structural break tests. We first utilized the Quandt-Andrews unknown breakpoints test for the period between the middle 1970s and late 1980s. The results showed no evidence of a break. However, when we tested for structural breaks using the Chow break test for indi-

vidual quarters starting from the middle 1970s, both models reveal structural break points. Furthermore, the break points are spread out. In model 2, we found that in most quarters starting in 1978Q3, the Chow tests show significant statistics for a break at conventional levels. Interestingly, the F-statistics reach a maximum at 2000Q2. The results for model 3 reveal a significant break in each quarter between 1982Q1 and 1986Q2, at conventional levels. The F-statistics reach a maximum at 1983Q4. We interpret the results that break points are spread out over a period as a gradual structural change.¹³

[Table 2 about here.]

To clarify the structural change between the earlier and later periods, we divide the sample at the fourth quarter of 1982. This quarter is chosen because the NBER reports November 1982 as the trough of the business cycle period of July 1981-November 1982. The Chow break test results for 1982Q4 and 1983Q1 for models 2 and 3 are reported in table 2.

[Table 3 about here.]

Table 3 reports results examining the unit root properties of the variables in this divided sample. We tested the variables for unit roots using ADF test statistics. We specified a constant and linear time trend, and lag lengths were determined by the SIC. For the later cycles, we include the financial obligation and debt service ratios in our tests.

We observe interesting differences between the earlier and later periods. We find that, for the later cycles, all variables have a unit root. However, in the earlier cycles, we observe that none of the household variables (i.e., net worth, household debt, consumer debt, and mortgage debt) show evidence of a unit root. This finding is also evidence that structural changes occurred in U.S. economy, particularly in the household sector.

2.2 Estimation Results for Sub-periods

Table 4 reports the results for the earlier period (1951Q4-1982Q4), while table 5 reports the regression results for the later period (1983Q1-2009Q1). Regression specifications presented

in models 2 and 3 from table 1, which were estimated for the entire period, are also estimated for the separated samples in tables 4 and 5. The results from model 2 in table 4 show that the changes in household debt and net worth, and the level of net worth are significant with the expected positive signs, but the level of household debt is not significant in the earlier period. However, in the later period (table 5), all the household debt variables are significant with the expected signs. The household debt level has a negative coefficient, while all the other variables are positive and significant.

[Table 4 about here.]

[Table 5 about here.]

In model 3 for both periods, we observe that the coefficients for many of the household debt variables lose significance. In the later period, only the change in household debt and the level of net worth are significant, and only the change in net worth is significant for the earlier period regression.¹⁴ In model 3, the level of investment has a significant negative coefficient for the earlier period, but is not significant for the later period. We also observe that introducing the level and first difference of investment into the regression corrects the serial correlation problem in the later period, as in the regression for the whole period.

For the later period, we also control for two household debt burden measures which are only available for the later period: household financial obligations as a percent of disposable personal income and household debt service payments as a percent of disposable personal income. Table 5 reports the results from models that incorporate these measures of the debt burden. Models 4 and 5 are the same as model 2 with the debt burden measures included. We observe that all household debt variables, except the change in net worth, are significant with the expected signs—including the negative coefficient on the level of household debt. The debt service burdens, in both level and change, have negative effects on output. They have significant negative coefficients with a large magnitude. The large magnitude is due to

the unit difference since the debt burden measures are in percent of disposable income, but all other variables are in terms of real dollars (millions).

We also controlled for the debt burden variables in model 3. These regressions are reported as models 6 and 7 in table 5. Only the change in household debt and the level of net worth have significant coefficients among the household and debt service burden variables.¹⁵

In summary, regression analysis for the earlier period indicates no evidence that the household debt variables had any negative effect on output. However, according to the analysis for the later period, there is some evidence that the accumulation of household debt has negative effects on output. We also see some evidence that new household borrowing could boost output for the both periods.¹⁶ Therefore, our results also provide a supporting evidence for debt-driven cycles as in Palley (1994).¹⁷

3 Conclusion

Our empirical results indicate evidence of a structural change in the relationship of U.S. output to household debt. The Chow break tests indicate multiple breaks points, which we interpret as evidence for a gradual structural change due to financialization. Unit root tests for the separate samples showed that the data generating processes of the household variables are fundamentally different between the earlier and later periods. None of the household variables possesses a unit root in the earlier period, yet all of them do in the later period.

The ADL regression analysis for the whole sample period indicates that household financial variables in general have effects on output—including a negative effect of the level of household debt, as hypothesized. To account for structural change during the period of financial liberalization, we have divided the sample at the fourth quarter of 1982. In the ADL regression analysis for the earlier period, we found no evidence that the household debt variables had any negative effect on output. However, according to the regression analysis

for the later period, there is evidence that the household indebtedness has negative effects on output. Our results suggest structural differences between the earlier and later periods in the effect of household debt on the U.S. macroeconomy.

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Notes

¹ These two series are available starting in 1980. The debt service ratio measures the share of income committed by households for paying interest and principal on their debt. The financial obligations ratio, in addition to including debt payments, incorporates households' other recurring expenses—such as rents, auto leases, homeowners' insurance and property taxes—that may be subtracted from the uncommitted income available to households (Greenspan, 2004).

²The financial instability hypothesis originally emphasize firms's investment financing behavior. Minsky argues that a prolonged period of prosperity will induce euphoric expectations, leading firms to adopt riskier and riskier financial stances. As the average firm evolves from “hedge” to largely “speculative” and even “Ponzi” finance, the economy becomes systematically fragile and susceptible to a sudden financial crisis (Minsky, 1986).

³There are few other empirical studies that provide rather inconsistent evidence. Garner (1996) and Schmitt (2000) find that some macroeconomic indicators (e.g., real GDP) predict the various consumer debt measures in the Granger causality sense, but not the reverse. The Granger approach tests whether past values of one variable can improve the prediction of the value of another variable. Regressors in Palley's regression are all past values (lagged). Palley and Schmitt therefore provide conflicting results.

⁴All the variables in Palley's empirical analysis are in real, per capita terms.

⁵We attempted to replicate Palley's regression and VAR results. Although we could not obtain the exact data set Palley used, our replication results are similar to Palley's results.

⁶The gross national product (GNP) was the main macroeconomic aggregate used in the U.S. at the time of Palley's study.

⁷All data are seasonally adjusted except household net worth. The Fed flow of funds does not have seasonally adjusted household net worth series.

⁸This specification can be interpreted as a variant of dynamic Ordinary Least Square (DOLS), suggested by Stock and Watson (1993) as a method that is robust to the inclusion of nonstationary and possibly cointegrated data. In the Stock-Watson DOLS method, the coefficients on the variables in levels can be interpreted as the long-run relationships.

⁹A similar empirical modeling strategy was adopted by Stockhammer (2004), who explores the linkage between financialization and capital accumulation.

¹⁰Our specification is somewhat different from Palley's specification. In Palley's work, the real debt burden is proxied by the level of real per capita consumer installment debt multiplied by the ex post real prime rate. Similar, but more broadly defined debt burden measures are incorporated in the section 2.2. Palley also incorporates a nominal debt burden measure, nominal prime interest rate, and a measure of inflation tax as regressors. We drop these variables since our variables are all real, and our focus is on the impact of the real value of debt accumulation on real GDP.

¹¹This may be due to the multicollinearity between investment and household debt, since household debt includes mortgage debt and private fixed investment includes residential construction.

¹²Palley (2007) defines financialization as “a process whereby financial markets, financial institutions, and financial elites gain greater influence over economic policy and economic outcomes.” Epstein (2005) similarly defines financialization as “the increasing role of financial motives, financial markets, financial actors and financial institutions in the operation of the domestic and international economies.”

¹³It is interesting to note that Stockhammer (2004) did not find structural break points in 1980 using the Chow break test in his work on financialization. His regression analysis is

on the linkage between a measure of financialization and capital accumulation.

¹⁴ For the later period, the coefficient on the level of household debt is sensitive to the inclusion of a constant in the regression. Without a constant, this variable has a significant negative coefficient as in model 2.

¹⁵ The coefficients on the level of household debt are again sensitive to the inclusion of a constant in the regressions. Without the constant, the variable has a significant negative coefficient in both models 6 and 7.

¹⁶ For a comparison, we also produced the regressions in per worker unit. They are reported in tables 6, 7, and 8. The results are largely similar. In the earlier period, there is no evidence of negative effect of household debt variables on output. For the later period, there is some evidence that the accumulation of household debt has negative effects on output. However the evidence is relatively weaker. The coefficient for the level of household debt per worker is negative and significant only for the model 2 in table 8. There is also some evidence of the positive effect of new household borrowing. However, the evidence is again weaker. The coefficients for the change in household debt per worker are positive and significant only for the models 3 and 4 for the later period in table 8.

¹⁷ Palley (1994) uses the term a Kaldor-Minsky cycle for a debt-driven cycle. A rise in household debt initially increases household expenditure and hence promotes growth, but eventually the accumulation of debt becomes excessive. This implies that there is a transfer of income from low saving agents (debtors) to high saving agents (rentiers) at an increasing rate due to the debt service payments. The debt service burden then reduces household expenditure and output. This could provide a mechanism of a credit-driven cyclical process of output fluctuations.

A Data: Sources and Definitions

Sources:

BEA: Bureau of Economic Analysis

<http://www.bea.gov>

FED Flow of Fund: Federal Reserve Board Flow of Fund

<http://www.federalreserve.gov/releases/z1/Current/data.htm>

BLS: Bureau of Labor Statistics

<http://www.bls.gov>

| Variables | Source |
|--------------------------------|------------------|
| Real GDP | BEA |
| Chain-type Price Index for PCE | BEA |
| Consumer Debt | FED Flow of Fund |
| Household Debt | FED Flow of Fund |
| Investment | BEA |
| FODSP | FED Flow of Fund |
| TDSP | FED Flow of Fund |
| Labor Force | BLS |

PCE: personal consumption expenditures

FODSP: household financial obligations as a percent of disposable personal income

TDSP: household debt service payments as a percent of disposable personal income

Consumer debt is households and nonprofit organizations consumer credit liability from Federal Reserve statistical release Z.1, FED Flow of Funds. The identification number is Z1/Z1/LA153166000.Q for the seasonally adjusted.

Household debt is households and nonprofit organizations credit and equity market instruments liability from Federal Reserve statistical release Z.1, FED Flow of Funds. The identification number is Z1/Z1/LA154102005.Q for the seasonally adjusted.

Household net worth is households and nonprofit organizations net worth (market value) asset from Federal Reserve statistical release Z.1, Fed Flow of Funds. The identification number is Z.1:FL152090005.Q.

FODSP is seasonally adjusted and from Fed Flow of Fund. The identification number is FOR/FOR/DTFDpercentYPD.Q.

TDSP is seasonally adjusted and from Fed Flow of Fund. The identification number is FOR/FOR/DTFpercentYPD.Q.

Investment is the seasonally adjusted real fixed private investment data from the Bureau of Economic Analysis.

Labor force is the seasonally adjusted monthly total labor force from the Bureau of Labor Statistics. The last month of the each quarter is used to match with the other quarterly data series.

B Figures and Tables

B.1 figures

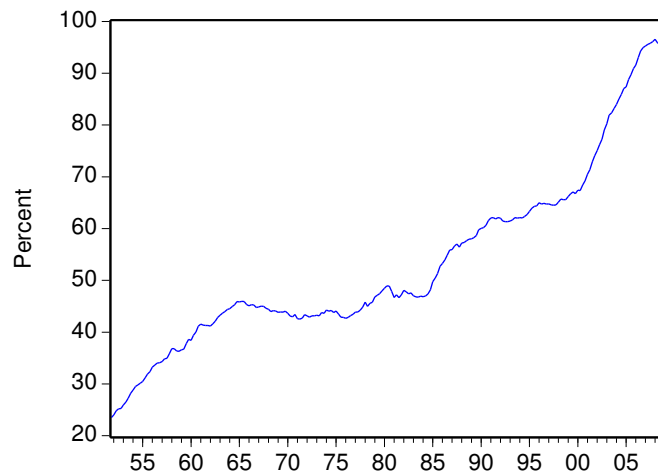


Figure 1: Household Debt-GDP Ratio (1951Q4-2009Q1)

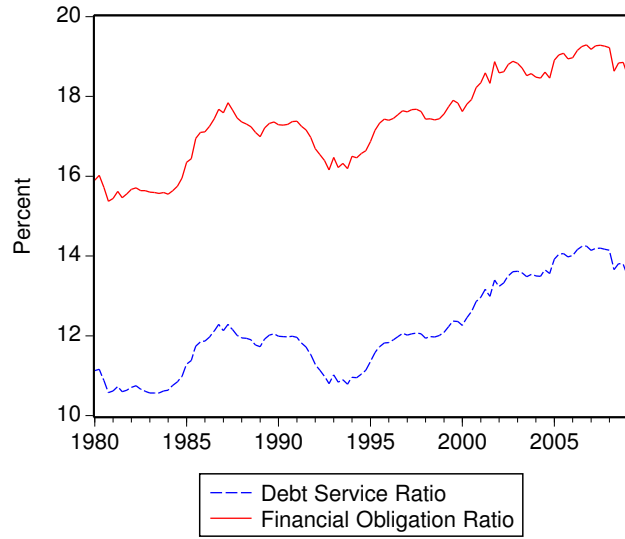


Figure 2: Debt Service and Financial Obligation Ratios (1980Q1-2009Q1)

B.2 Tables

Table 1: ADL Regressions: Sample period 1951Q4-2009Q1

| | Model1 | Model2 | Model3 |
|---|-----------------------|-----------------------|-----------------------|
| Constant | -14962.89 (-1.168) | -11145.54 (-0.948) | -23310.51* (-1.75) |
| Real GDP _{t-1} | 1.009*** (127.622) | 1.002*** (165.039) | 1.009*** (162.65) |
| Household debt _{t-1} | -0.02** (-2.555) | -0.03*** (-4.724) | -0.022*** (-3.477) |
| Change in household debt _{t-1} | 0.264*** (2.671) | 0.254*** (3.211) | 0.289*** (3.474) |
| Consumer debt _{t-1} | -0.084 (-1.561) | | |
| Change in consumer debt _{t-1} | 0.390 (1.254) | | |
| Net worth _{t-1} | 0.006*** (2.921) | 0.006*** (3.06) | 0.01*** (3.80) |
| Change in net worth _{t-1} | 0.014*** (2.776) | 0.014*** (2.864) | 0.003 (0.701) |
| Investment _{t-1} | | | -0.185*** (-3.318) |
| Change in investment _{t-1} | | | 1.038*** (6.342) |
| Adjusted R-squared | 0.999 | 0.999 | 0.999 |
| Godfrey-Breusch LM (2) | 3.115 | 4.264 | 2.316 |
| | 0.046 | 0.015 | 0.101 |
| ARCH(1) | 1.998 | 1.305 | 1.152 |
| | 0.159 | 0.255 | 0.284 |

*, **, and *** denote significance at 10, 5 and 1 percent levels, respectively. *t*-statistics in parentheses.

Notes: Dependent variable: real GDP. Figures for the LM and ARCH tests are *F*-statistics with *p*-values.

Table 2: Chow structural break tests: 1951Q4-2009Q1

| Equation | Time | F-statistics | p-value |
|----------|--------|--------------|---------|
| Model 2 | 1982Q4 | 3.301 | 0.0067 |
| | 1983Q1 | 3.423 | 0.0053 |
| Model 3 | 1982Q4 | 2.596 | 0.0136 |
| | 1983Q1 | 2.600 | 0.0135 |

Notes: The null hypothesis is that there is no structural break.

Table 3: Unit roots tests for the separated samples

| | 1951Q4-1982Q4 | 1983Q1-2009Q1 |
|----------------------------|---------------|---------------|
| Real GDP | -2.547 | -1.628 |
| | 0.305 | 0.775 |
| Household debt | -4.198*** | -1.678 |
| | 0.006 | 0.754 |
| Mortgage debt | -3.938** | -1.713 |
| | 0.013 | 0.739 |
| Consumer debt | -4.131*** | -1.898 |
| | 0.008 | 0.648 |
| Net worth | -3.467** | -2.741 |
| | 0.048 | 0.223 |
| Investment | -2.962 | 0.002 |
| | 0.147 | 0.996 |
| Financial obligation ratio | | -2.255 |
| | | 0.453 |
| Debt service ratio | | -1.815 |
| | | 0.69 |

*, **, and *** denote significance at 10, 5 and 1 percent respectively. ADF test statistics with p -values are reported .

Tests are based on ADF statistic. The null hypothesis is that the variable has a unit root.

Table 4: ADL Regression: 1951Q4-1982Q4

| | Model2 | Model3 |
|--|-----------------------|-----------------------|
| Constant | -4068.984 (-0.099) | -3682.116 (-0.077) |
| Real GDP _{<i>t</i>-1} | 0.941*** (31.201) | 0.987*** (25.699) |
| Household debt _{<i>t</i>-1} | 0.005 (0.098) | 0.041 (0.740) |
| Change in household debt _{<i>t</i>-1} | 0.785*** (3.586) | 0.416 (1.346) |
| Net worth _{<i>t</i>-1} | 0.018* (1.757) | 0.01 (1.016) |
| Change in net worth _{<i>t</i>-1} | 0.045** (2.609) | 0.045*** (2.629) |
| Investment _{<i>t</i>-1} | | -0.302** (-1.900) |
| Change in Investment _{<i>t</i>-1} | | 1.012*** (2.779) |
| Adjusted R-squared | 0.999 | 0.998 |
| LM (2) | 0.339 | 1.235 |
| | 0.714 | 0.295 |
| ARCH(1) | 0.257 | 0.450 |
| | 0.613 | 0.504 |

*, **, and *** denote significance at 10, 5 and 1 percent, respectively. T-statistics in parentheses.
Notes: Dependent variable: real GDP. Figures for the LM and ARCH tests are *F*-statistics with *p*-values.

Table 5: ADL regression: 1983Q1-2009Q1

| | Model2 | Model3 | Model4 | Model5 | Model6 | Model7 |
|---|-----------|----------|-------------|-------------|-----------|-----------|
| Constant | 97398.04* | 86350.82 | 379828.7** | 377247.9*** | 74652.86 | 150753.9 |
| | (1.830) | (1.351) | (2.201) | (2.636) | (0.411) | (0.916) |
| Real GDP _{t-1} | 0.978*** | 0.982*** | 0.980*** | 0.974*** | 0.978*** | 0.977*** |
| | (67.590) | (57.787) | (69.805) | (68.439) | (56.809) | (54.484) |
| Household debt _{t-1} | -0.023*** | -0.013 | -0.021*** | -0.016* | -0.013 | -0.012 |
| | (-2.837) | (-1.425) | (-2.639) | (-1.692) | (-1.467) | (-1.159) |
| Change in household debt _{t-1} | 0.214** | 0.223** | 0.307*** | 0.359*** | 0.238** | 0.257** |
| | (2.344) | (2.262) | (3.275) | (3.611) | (2.327) | (2.472) |
| Net worth _{t-1} | 0.008*** | 0.008** | 0.009*** | 0.009*** | 0.008** | 0.007** |
| | (3.251) | (2.563) | (3.518) | (3.637) | (2.359) | (2.273) |
| Change in net worth _{t-1} | 0.011* | 0.002 | 0.008 | 0.007 | 0.002 | 0.002 |
| | (1.899) | (0.425) | (1.497) | (1.309) | (0.430) | (0.448) |
| Investment _{t-1} | | -0.067 | | | -0.038 | -0.022 |
| | | (-0.833) | | | (-0.456) | (-0.249) |
| Change in Investment _{t-1} | | 1.023*** | | | 0.966*** | 0.926*** |
| | | (4.915) | | | (4.054) | (3.894) |
| Financial obligation ratio _{t-1} | | | -18808.21* | | 1798.180 | |
| | | | (-1.703) | | (0.156) | |
| Change in financial obligation ratio _{t-1} | | | -65182.30** | | -42228.64 | |
| | | | (-2.190) | | (-1.452) | |
| Debt service ratio _{t-1} | | | | -26613.07** | | -4438.265 |
| | | | | (-2.151) | | (-0.335) |
| Change in debt service ratio _{t-1} | | | | -79074.87** | | -57896.48 |
| | | | | (-2.019) | | (-1.493) |
| Adjusted R-squared | 0.999 | 0.999 | 0.999 | 0.999 | 0.999 | 0.999 |
| LM (2) | 3.955 | 2.171 | 2.656 | 2.468 | 2.789 | 2.829 |
| | 0.022 | 0.120 | 0.075 | 0.090 | 0.067 | 0.064 |
| ARCH(1) | 2.321 | 0.917 | 2.449 | 1.989 | 0.435 | 0.645 |
| | 0.131 | 0.341 | 0.121 | 0.162 | 0.511 | 0.424 |

*, **, and *** denote significance at 10, 5 and 1 percent, respectively. T-statistics in parentheses.

Notes: Dependent variable: real GDP. Figures for the LM and ARCH tests are F -statistics with p -values.

Table 6: ADL Regressions (per worker unit): 1951Q4-2009Q1

| | Model1 | Model2 | Model3 |
|---|----------------------|----------------------|----------------------|
| Constant | -0.219 (-0.517) | -0.107 (-0.322) | -0.361 (-0.783) |
| Real GDP _{t-1} | 0.999*** (62.262) | 0.993*** (91.539) | 1.004*** (70.441) |
| Household debt _{t-1} | -0.014 (-1.304) | -0.020** (-2.034) | -0.012 (-1.188) |
| Change in household debt _{t-1} | 0.140 (0.984) | 0.170 (1.615) | 0.134 (1.164) |
| Consumer debt _{t-1} | -0.056 (-0.744) | | |
| Change in consumer debt _{t-1} | 0.306 (0.743) | | |
| Net worth _{t-1} | 0.006** (2.171) | 0.006** (2.255) | 0.009** (2.519) |
| Change in net worth _{t-1} | 0.011 (1.632) | 0.011* (1.653) | 0.003 (0.489) |
| Investment _{t-1} | | | -0.158** (-2.006) |
| Change in investment _{t-1} | | | 0.990*** (4.518) |
| Adjusted R-squared | 0.998 | 0.998 | 0.998 |
| Godfrey-Breusch LM (2) | 0.628 | 1.036 | 2.584 |
| ARCH(1) | 0.534 | 0.356 | 0.077 |
| | 4.179 | 3.460 | 6.520 |
| | 0.042 | 0.064 | 0.011 |

*, **, and *** denote significance at 10, 5 and 1 percent levels, respectively. *t*-statistics in parentheses.

Notes: Dependent variable: real GDP. Figures for the LM and ARCH tests are *F*-statistics with *p*-values.

Table 7: ADL Regression (per worker unit): 1951Q4-1982Q4

| | Model2 | Model3 |
|---|----------------------|----------------------|
| Constant | 0.789 (0.777) | 1.852 (1.600) |
| Real GDP _{t-1} | 0.907*** (24.950) | 0.912*** (18.503) |
| Household debt _{t-1} | 0.033 (0.670) | 0.084* (1.664) |
| Change in household debt _{t-1} | 0.193 (0.790) | -0.360 (-1.147) |
| Net worth _{t-1} | 0.019* (1.906) | 0.008 (0.866) |
| Change in net worth _{t-1} | 0.039** (2.180) | 0.046** (2.572) |
| Investment _{t-1} | | -0.121 (-0.731) |
| Change in Investment _{t-1} | | 1.282*** (3.024) |
| Adjusted R-squared | 0.993 | 0.994 |
| LM (2) | 0.433 | 0.800 |
| | 0.649 | 0.451 |
| ARCH(1) | 0.304 | 1.099 |
| | 0.582 | 0.296 |

*, **, and *** denote significance at 10, 5 and 1 percent, respectively. T-statistics in parentheses.
Notes: Dependent variable: real GDP. Figures for the LM and ARCH tests are F -statistics with p -values.

Table 8: ADL regression (per worker unit): 1983Q1-2009Q1

| | Model2 | Model3 | Model4 | Model5 | Model6 | Model7 |
|---|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| Constant | 1.286 (1.544) | 1.926* (1.810) | 3.692** (2.364) | 3.582** (2.539) | 1.895 (1.074) | 2.791 (1.550) |
| Real GDP _{t-1} | 0.966*** (41.702) | 0.950*** (33.041) | 0.970*** (42.853) | 0.960*** (41.671) | 0.945*** (32.457) | 0.941*** (31.038) |
| Household debt _{t-1} | -0.017* (-1.758) | -0.001 (-0.040) | -0.013 (-1.336) | -0.006 (-0.575) | -0.001 (-0.094) | 0.002 (0.157) |
| Change in household debt _{t-1} | 0.158 (1.470) | 0.109 (0.933) | 0.263** (2.343) | 0.322*** (2.710) | 0.133 (1.091) | 0.165 (1.330) |
| Net worth _{t-1} | 0.007** (2.504) | -0.004 (1.460) | 0.008*** (2.762) | 0.009*** (2.925) | 0.005 (1.290) | 0.004 (1.214) |
| Change in net worth _{t-1} | 0.002 (0.321) | -0.004 (-0.644) | 0.263 (-0.090) | -0.002 (-0.308) | -0.004 (-0.651) | -0.004 (-0.638) |
| Investment _{t-1} | | 0.025 (0.255) | | | 0.053 (0.524) | 0.081 (0.762) |
| Change in Investment _{t-1} | | 1.103*** (4.489) | | | 1.051*** (3.614) | 0.965*** (3.335) |
| Financial obligation ratio _{t-1} | | | -0.173* (-1.821) | | 0.014 (0.140) | |
| Change in financial obligation ratio _{t-1} | | | -0.446* (-1.743) | | -0.319 (-1.268) | |
| Debt service ratio _{t-1} | | | | -0.227** (-2.132) | | -0.058 (-0.488) |
| Change in debt service ratio _{t-1} | | | | -0.576* (-1.697) | | -0.504 (-1.499) |
| Adjusted R-squared | 0.998 | 0.998 | 0.998 | 0.998 | 0.998 | 0.998 |
| LM (2) | 0.823 | 4.395 | 0.758 | 0.841 | 5.585 | 5.815 |
| ARCH(1) | 0.442 | 0.014 | 0.471 | 0.434 | 0.005 | 0.004 |
| | 4.842 | 3.967 | 4.630 | 5.428 | 3.361 | 3.578 |
| | 0.030 | 0.049 | 0.033 | 0.021 | 0.069 | 0.061 |

*, **, and *** denote significance at 10, 5 and 1 percent, respectively. T-statistics in parentheses.

Notes: Dependent variable: real GDP. Figures for the LM and ARCH tests are F -statistics with p -values.

